

# Asymptotic Consistency and Inconsistency of the Chain Ladder

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# Overview

- Claims reserving in non-life insurance
- Chain ladder model
- Open problem: consistency of the development factors
- Simulations and real data example

- Based on:

Pešta and Hudecová (2012). Asymptotic consistency and inconsistency of the chain ladder. *Insurance: Mathematics and Economics*, 51(2): 472–479.

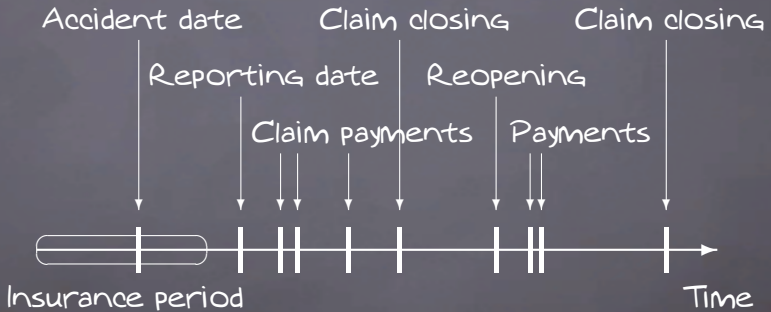
- Support:

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# Non-life insurance

- Operates on the lines of Business (LoB):
  - ▶ motor/car insurance (motor third party liability, motor hull)
  - ▶ property insurance (private and commercial insurance against fire, water, flooding, Business interruption, ...)
  - ▶ liability insurance
  - ▶ accident insurance
  - ▶ health insurance
  - ▶ marine insurance (including transportation)
  - ▶ other (aviation, travel insurance, legal protection, credit insurance, epidemic insurance, ...)
- Life insurance products are rather different, e.g., terms of contracts, type of claims, risk drivers

# Timeline of a claim



# Settlement of a claim

- Reporting delay (Between occurrence and reporting) – can take several years (liability insurance: asbestos or environmental pollution claims)
- After being reported to the insurer – several years may elapse before the claim is finally settled (fast in property insurance, liability or bodily injury claims: long time before the total circumstances are clear and known)
- Reopening – (unexpected) new developments, or if a relapse occurs

# Reserving

- Claims reserves represent the money which should be held by the insurer so as to be able to meet all future claims arising from policies currently in force and policies written in the past
- Most non-life insurance contracts are written for a period of one year
- Only one payment of premium at the start of the contract in exchange for coverage over the year
- Reserves are calculated by forecasting future losses from past losses

# Terminology

- $X_{i,j}$  ... claim amounts in development year  $j$  with accident year  $i$
- $X_{i,j}$  stands for the incremental claims in accident year  $i$  made in accounting year  $i + j$
- $n$  ... current year - corresponds to the most recent accident year and development period
- Our data history consists of right-angled isosceles triangles  $X_{i,j}$ , where  $i = 1, \dots, n$  and  $j = 1, \dots, n + 1 - i$

# Notation

- $C_{i,j}$  ... cumulative payments in origin year  $i$  after  $j$  development periods

$$C_{i,j} = \sum_{k=1}^j X_{i,k}$$

- $C_{i,j}$  ... a random variable of which we have an observation if  $i + j < n + 1$
- Aim is to estimate the ultimate claims amount  $C_{i,n}$  and the outstanding claims reserve

$$R_i = C_{i,n} - C_{i,n+1-i}, \quad i = 2, \dots, n$$



# Run-off triangle

Accident year $i$	Development year $j$				
	1	2	...	$n-1$	$n$
1	$C_{1,1}$	$C_{1,2}$	...	$C_{1,n-1}$	$C_{1,n}$
2	$C_{2,1}$	$C_{2,2}$	...	$C_{2,n-1}$	
$\vdots$	$\vdots$	$\vdots$	$\ddots$		
$\vdots$	$\vdots$	$\vdots$	$C_{i,n+1-i}$		
$n-1$	$C_{n-1,1}$	$C_{n-1,2}$			
$n$	$C_{n,1}$				

# Chain ladder

Mack (1993)

- [1]  $E[C_{i,j+1}|C_{i,1}, \dots, C_{i,j}] = f_j C_{i,j}$
- [2]  $\text{Var}[C_{i,j+1}|C_{i,1}, \dots, C_{i,j}] = \sigma_j^2 C_{i,j}$
- [3] Accident years  $[C_{i,1}, \dots, C_{i,n}]$  are independent vectors

# Development factors $f_j$

$$\widehat{f}_j^{(n)} = \frac{\sum_{i=1}^{n-j} C_{i,j+1}}{\sum_{i=1}^{n-j} C_{i,j}}, \quad 1 \leq j \leq n-1$$

$$\widehat{f}_n^{(n)} \equiv 1 \quad (\text{assuming no tail})$$

# Properties

- Ultimate claims amounts  $C_{i,n}$  are estimated by

$$\widehat{C}_{i,n} = C_{i,n+1-i} \times \widehat{f}_{n+1-i}^{(n)} \times \cdots \times \widehat{f}_{n-1}^{(n)}$$

- Under the assumptions [1], [3], and

$$[4] \quad \sum_{i=1}^{n-j} C_{i,j} > 0$$

$\widehat{f}_j^{(n)}$  are unbiased and mutually uncorrelated

- Assumption [2] is essential for the standard error of  $\widehat{C}_{i,n}$

# Open problem

- From some point of view, consistency is more important than unbiasedness
- E.g., Bornhuetter-Ferguson method uses

$$\widehat{\beta}_j^{(n)} = \prod_{k=j}^{n-1} \frac{1}{\widehat{f}_k^{(n)}}$$

- The unbiasedness of  $\widehat{f}_j^{(n)}$  does not "transfer" to  $\widehat{\beta}_j^{(n)}$  in any sense
- Ex:  $Y_1, \dots, Y_n$  iid with finite  $EY$
- $T_1(Y_1, \dots, Y_n) = Y_1$  vs  $T_2(Y_1, \dots, Y_n) = \frac{1}{n} \sum_{i=1}^n Y_i + \frac{1}{n}$

# Conditional convergence

$\xi_n \xrightarrow[n \rightarrow \infty]{[P_\zeta]\text{-a.s.}} \chi$ ,  $[P]$ -a.s. means

$$P \left[ P_\zeta \left\{ \lim_{n \rightarrow \infty} \xi_n = \chi \right\} = 1 \right] = 1$$

$\xi_n \xrightarrow[n \rightarrow \infty]{P_{\zeta_n}} \chi$ ,  $[P]$ -a.s. means

$$\forall \varepsilon > 0 : P \left[ \lim_{n \rightarrow \infty} P_{\zeta_n} \{ |\xi_n - \chi| \geq \varepsilon \} = 0 \right] = 1$$

$\xi_n \xrightarrow[n \rightarrow \infty]{L_p(P_{\zeta_n})} \chi$ ,  $[P]$ -a.s. ( $p \geq 1$ ) means

$$P \left[ \lim_{n \rightarrow \infty} E_{\zeta_n} |\xi_n - \chi|^p = 0 \right] = 1$$

# Conditioning

- Conditional convergence in probability and in  $L_p$  along some sequence of random variables  $\{\zeta_n\}_{n=1}^{\infty}$  can be defined, because the concept of these two types of convergence comes from a topology
- Despite of that, the almost sure convergence does not correspond to a convergence with respect to any topology and, hence, it is not metrizable
- Thereafter, the conditional convergence almost surely cannot be defined along a sequence of random variables, but only given one random variable  $\zeta$

# Consistency

Denote  $D_j^{(n)} = \{C_{i,k} : k \leq j, i \leq n - j + 1\}$  and  $D_j = \{C_{i,k} : k \leq j, i \in \mathbb{N}\}$ . Then (i)–(iv) are equivalent:

(i)

$$\widehat{f}_j^{(n)} \xrightarrow[n \rightarrow \infty]{[P_{D_j}]^{-a.s.}} f_j, \quad [P]^{-a.s.};$$

(ii)

$$\widehat{f}_j^{(n)} \xrightarrow[n \rightarrow \infty]{P_{D_j^{(n)}}} f_j, \quad [P]^{-a.s.};$$

(iii)

$$\widehat{f}_j^{(n)} \xrightarrow[n \rightarrow \infty]{L_2(P_{D_j^{(n)}})} f_j, \quad [P]^{-a.s.};$$

(iv)

$$\sum_{i=1}^{n-j} C_{i,j} \xrightarrow[n \rightarrow \infty]{} \infty, \quad [P]^{-a.s.}$$



# Remark 1

Due to the independence of the different accident years (assumption [3]), the statements (ii) and (iii) can be equivalently replaced by

$$\widehat{f}_j^{(n)} \xrightarrow[n \rightarrow \infty]{P_{D_j}} f_j, [P]-a.s.$$

and

$$\widehat{f}_j^{(n)} \xrightarrow[n \rightarrow \infty]{L_2(P_{D_j})} f_j, [P]-a.s.,$$

respectively.

# Remark II

- Unconditional consistency in case of the  $L_2$  convergence
- $\widehat{f}_j^{(n)} \rightarrow f_j$  in  $L_2$  (unconditionally) as  $n \rightarrow \infty$  iff

$$E \left[ \frac{1}{\sum_{i=1}^{n-j} C_{i,j}} \right] \rightarrow 0, \quad n \rightarrow \infty$$

- This condition is obviously more complicated than the condition (iv), and it is practically unverifiable
- Thus, the conditional convergence is not only more natural one in this case, but even more convenient one

# Rate of convergence

- Consistency of an estimator is a very important but only qualitative property
- Measure consistency – quantitative way
- Denote the conditional mean square error of the estimate of development factor  $f_j$  as

$$MSE \left( \hat{f}_j^{(n)} \right) := E \left\{ \left[ \hat{f}_j^{(n)} - E \left( \hat{f}_j^{(n)} \right) \right]^2 \mid D_j^{(n)} \right\}.$$

Then, with probability one holds

$$MSE \left( \hat{f}_j^{(n)} \right) = \mathcal{O} \left( \left[ \sum_{i=1}^{n-j} C_{i,j} \right]^{-1} \right), \quad n \rightarrow \infty.$$

# On the rate of convergence

- Complete characterization of the conditional convergence of development factors' estimate
- The slower (faster) divergence of

$$\sum_{i=1}^{\infty} C_{i,j}$$

implies the slower (faster) realization of consistency of the development factors' estimates

# On the necessary and sufficient condition

Let  $j \in \mathbb{N}$  be fixed. Then the following conditions are equivalent.

1. The condition (iv) holds.

2.

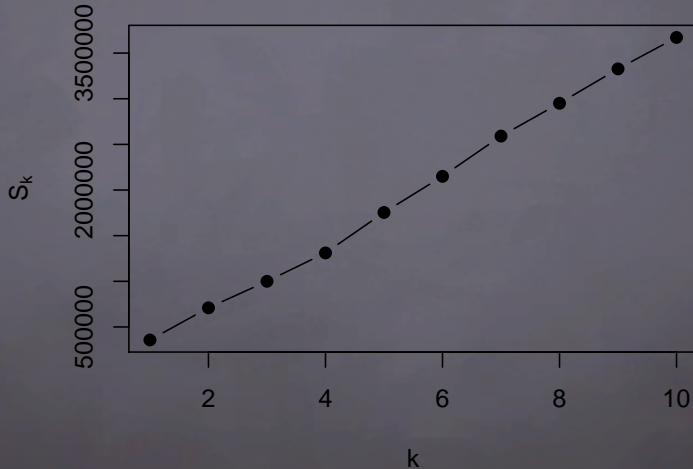
$$\sum_{i=1}^{\infty} EC_{i,1} = \infty. \quad (1)$$

3. The condition (iv) holds for  $j_0 \in \mathbb{N}$ ,  $j \neq j_0$ .

# Practical aspects

- Either  $\hat{f}_j^{(n)}$  is consistent for  $f_j$  for all  $j \in \mathbb{N}$ , or none of them is consistent
- The consistency of  $\hat{f}_j^{(n)}$  is equivalent to the condition  $\sum_{i=1}^n C_{i,1} \rightarrow \infty$ , [P]-a.s. as  $n \rightarrow \infty$
- Denote  $S_k = \sum_{i=1}^k C_{i,1}$ ,  $k = 1, \dots, n$  the cumulative sums of the cumulative claims  $C_{i,1}$  in the first development year
- For instance, the ratios  $C_{k+1,1}/C_{k,1}$  or the sequence  $\sqrt[k]{C_{k,1}}$  can be studied
- Artificial data set Taylor and Ashe (1983)

# Real data example



# Inconsistency

- What kinds of Business Behavior corresponds to the violation of consistency?
- For instance, condition (iv) can be violated if one observes a decreasing trend (decreasing fast enough) in payments across the accident years
- So to speak, the corresponding line of business is worsening maybe due to new insurance companies entering the market or changing (decreasing) prices of such insurance product
- Furthermore, splitting one existing line of business into several others can also cause inconsistency in the estimation of development factors



# Simulations

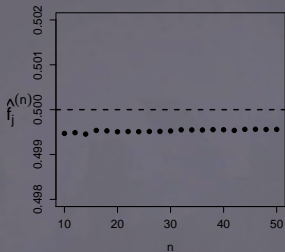
- $C_{i,1}$  was generated such that  $C_{i,1} \in L_2$  and  $C_{i,j} \geq 0$
- $f_j$  was set to  $f_j = j/(j+1)$
- $C_{i,2}, \dots, C_{i,N}$  were generated successively such that  $C_{i,j}$  satisfies [1] and [2]
- Hence, for each  $j$ ,  $C_{i,j}$  is drawn from a distribution with mean  $f_{j-1}C_{i,j-1}$  and variance  $\sigma_j^2 C_{i,j-1}$  for some  $\sigma_j^2 \in (0, \infty)$
- Since the accident years are assumed to be independent (assumption [3]), the rows of the data sets were generated separately, using the same approach
- $C_{i,1}$  were drawn from the exponential distribution, and the  $C_{i,j}$  was generated from the Poisson distribution with the parameter  $f_{j-1}C_{i,j-1}$  for  $j = 2, \dots, N$

# Decreasing Business

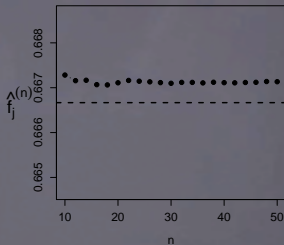
- consider a fast decreasing business, i.e., the situation where condition (I) does not hold
- $C_{i,1}$  was generated from the exponential distribution with the mean  $i^{-2} \times 10^6$
- $\hat{f}_j^{(n)}$  do not converge to the true value  $f_j$
- Their values are close to  $f_j$  in this setting, but the estimates are indeed not consistent
- The same simulations were run also for  $C_{i,1}$  with the uniform distribution: the differences between values of estimates  $\hat{f}_j^{(n)}$  and the true values  $f_j$  are more noticeable

# Example 1

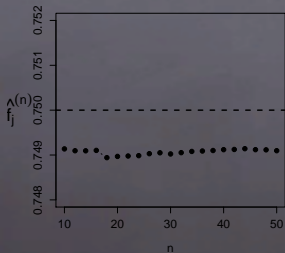
j = 1



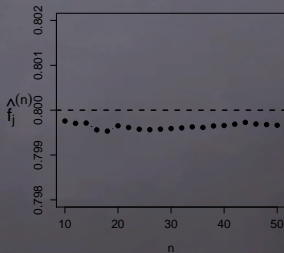
j = 2



j = 3



j = 4

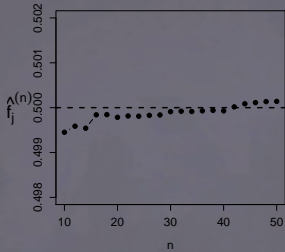


# Slowly decreasing Business

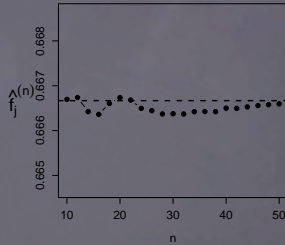
- situation where  $EC_{i,1}$  decreases with increasing  $i$ , but in a slow manner such that condition (I) holds
- $EC_{i,1} = i^{-1/2} \times 10^6$  in the exponential distribution
- clear convergence pattern can be observed, confirming that the estimates  $\hat{f}_j^{(n)}$  are consistent in this case

# Example II

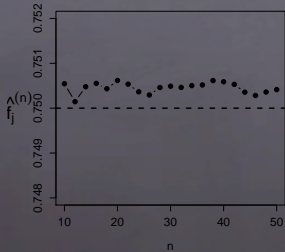
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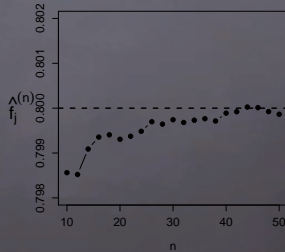
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j = 3



j = 4

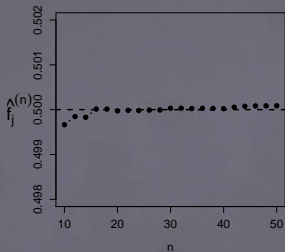


# Growing Business

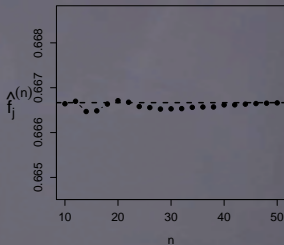
- $EC_{i,1}$  increases with increasing  $i$
- The parameters of the exponential distribution were set such that  $EC_{i,1} = \sqrt{i} \times 10^6$
- The figure obviously confirms that the estimates are consistent
- Moreover,  $\hat{f}_j^{(n)}$  converges to the true values  $f_j$  much faster than in previous case

# Example III

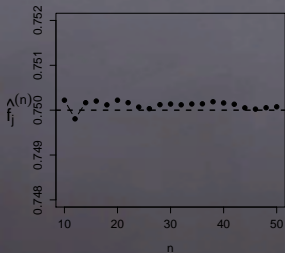
j = 1



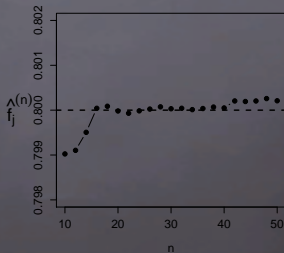
j = 2



j = 3



j = 4



# Conclusions

- conditional consistency and inconsistency of the development factors' estimate in the distribution-free chain ladder is investigated
- necessary and sufficient condition is derived
- weak, strong consistency, and consistency in the mean square are equivalent
- convergence rate is provided
- practical recommendations, how to check this necessary and sufficient condition, are discussed
- real data example and numerical simulations to illustrate the performance of the estimates
- possible violation of the condition with the consequences is demonstrated



Thank you !

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